

Linear estimates and LS-means in the doBy package

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1 Introduction

1.1 Linear functions of parameters

A linear function of a p -dimensional parameter vector β has the form

$$C = L\beta$$

where L is a $q \times p$ matrix which we call the *Linear Estimate Matrix* or simply *LE-matrix*. The corresponding linear estimate is $\hat{C} = L\hat{\beta}$. A *linear hypothesis* has the form $H_0 : L\beta = m$ for some q dimensional vector m .

1.2 LSmeans (population means, marginal means)

A special type of linear estimates is the so called *least-squares means* (or *LS-means*). Other names for these estimates include *population means* and *marginal means*. Consider an imaginary field experiment analyzed with model of the form

```
> lm( y ~ treat + block + year)
```

where **treat** is a treatment factor, **block** is a blocking factor and **year** is the year (a factor) where the experiment is repeated over several years. This model specifies the conditional mean $\mathbb{E}(Y|\text{treat}, \text{block}, \text{year})$. One may be interested in predictions of the form $\mathbb{E}(Y|\text{treat})$. This quantity can not formally be found from the model. However, it is tempting to average the fitted values of $\mathbb{E}(Y|\text{treat}, \text{block}, \text{year})$ across the levels of **block** and **year** and think of this average as $\mathbb{E}(Y|\text{treat})$. This average is precisely what is called the LS-means. If the experiment is balanced then this average is identical to the average of the observations when stratified according to **treat**.

An alternative is to think of **block** and **year** as random effects, for example:

```
> library(lme4)
> lmer( y ~ treat + (1|block) + (1|year))
```

In this case one would directly obtain $\mathbb{E}(Y|\text{treat})$ from the model. However, there are at least two reasons why one may be hesitant to consider such a random effects model.

- Suppose there are three blocks and the experiment is repeated over three consecutive years. This means that the random effects are likely to be estimated with a large uncertainty (the estimates will have only two degrees of freedom).
- Furthermore, treating **block** and **year** as random effects means they should in principle come from a large population of possible blocks and years. This may or may not be reasonable for the blocks, but it is a questionable assumption for the years.

1.3 A simulated dataset

In the following sections we consider these data:

```
> library(doby)
> set.seed(141164)
> dd <- expand.grid(A=factor(1:3), B=factor(1:3), C=factor(1:2))
> dd$y <- rnorm(nrow(dd))
> dd$x <- rnorm(nrow(dd))^2
> dd$z <- rnorm(nrow(dd))
> head(dd, 10)
```

```
##      A B C      y      x      z
```

```
## 1  1 1 1  0.73516 1.26234 -0.8019
## 2  2 1 1  0.18606 0.02870 -0.6705
## 3  3 1 1 -0.10532 1.09523 -1.3990
## 4  1 2 1 -0.31093 0.02323 -1.2826
## 5  2 2 1 -0.96994 2.72459  1.3790
## 6  3 2 1  0.59921 1.09027  0.7911
## 7  1 3 1 -0.02469 0.35388 -0.6327
## 8  2 3 1  0.38602 0.95420  0.6125
## 9  3 3 1  0.83019 2.33996  0.8216
## 10 1 1 2 -1.19164 0.06174  0.8128
```

Consider the additive model

$$y_i = \beta_0 + \beta_{A(i)}^1 + \beta_{B(i)}^2 + \beta_{C(i)}^3 + \beta^4 x_i + e_i \quad (1)$$

where $e_i \sim N(0, \sigma^2)$. We fit this model:

```
> mm <- lm(y ~ A + B + C + x, data=dd)
> coef(mm)
```

```
## (Intercept)          A2          A3          B2          B3          C2          x
##      0.05759      -0.47363      -0.38526      -0.17882       0.27992      -0.37859      0.31206
```

Notice that the parameters corresponding to the factor levels A1, B1 and C2 are set to zero to ensure identifiability of the remaining parameters.

1.4 LSmeans: details

LSmeans, population means and marginal means are used synonymously in the literature. These quantities are a special kind of contrasts as defined in Section 1.1. LSmeans seems to be the most widely used term, so we shall adopt this terms here too.

The model (1) is a model for the conditional mean $\mathbb{E}(y|A, B, C, x)$. Sometimes one is interested in quantities like $\mathbb{E}(y|A)$. This quantity can not formally be found because of the presences of unless B , C and x . However, suppose that A is a treatment of main interest, B is a blocking factor, C represents days on which the experiment was carried out and x denotes e.g. temperature for the specific block on the specific date. Then it is tempting to fix x at its average value \bar{x} and average $\mathbb{E}(y|A, B, C, \bar{x})$ over B and C (average over block and day) and think of this average as $\mathbb{E}(y|A)$.

The population mean for $A = 1$ is

$$\beta^0 + \beta_{A1}^1 + \frac{1}{3}(\beta_{B1}^2 + \beta_{B2}^2 + \beta_{B3}^2) + \frac{1}{2}(\beta_{C1}^3 + \beta_{C2}^3) \quad (2)$$

Recall that the parameters corresponding to the factor levels A1, B1 and C2 are set to zero to ensure identifiability of the remaining parameters. Therefore we may also write the population mean for $A = 1$ as

$$\beta^0 + \frac{1}{3}(\beta_{B2}^2 + \beta_{B3}^2) + \frac{1}{2}(\beta_{C2}^3) + \beta^4 \bar{x}. \quad (3)$$

This quantity can be estimated as (if $\bar{x} = 1.242$) :

```
> w <- c(1, 0, 0, 1/3, 1/3, 1/2, 1.242)
> sum(coef(mm) * w)

## [1] 0.2896
```

We may find the population mean for all three levels of A as

```
> L <- matrix(c(1, 0, 0, 1/3, 1/3, 1/2, 1.242,
                1, 1, 0, 1/3, 1/3, 1/2, 1.242,
                1, 0, 1, 1/3, 1/3, 1/2, 1.242), nr=3, byrow=TRUE)
> L %*% coef(mm)

##           [,1]
## [1,]  0.28958
## [2,] -0.18405
## [3,] -0.09568
```

Notice that the matrix W is based on that the first level of A is set as the reference level. If the reference level is changed then so must W be.

1.5 Using LEmatrix and LSmeans

Writing the matrix W is somewhat tedious and hence error prone. In addition, there is a potential risk of getting the wrong answer if the the reference level of a factor has been changed. The LEmatrix function provides an automated way of generating such matrices. The above W matrix is constructed by

```
> L <- LE_matrix(mm, effect='A')

## List of 2
## $ new.fact.lev:List of 1
## ..$ A: chr [1:3] "1" "2" "3"
## $ grid.data : 'data.frame': 3 obs. of 1 variable:
## ..$ A: chr [1:3] "1" "2" "3"

> L

##      (Intercept) A2 A3      B2      B3 C2      x
## [1,]          1  0  0 0.3333 0.3333 0.5 1.242
## [2,]          1  1  0 0.3333 0.3333 0.5 1.242
## [3,]          1  0  1 0.3333 0.3333 0.5 1.242
```

The **effect** argument requires to calculate the population means for each level of *A* aggregating across the levels of the other variables in the data.

The LSmeans function is simply a wrapper around first a call to LEmatrix followed by a call to (by default) `linest()`:

```
> linest(mm, L=L)

##      A      x estimate std.error statistic df p.value      lwr      upr
## 1 1 1.24    0.2895     0.318     0.910 11   0.382 -0.411 0.990
## 2 2 1.24   -0.1841     0.319    -0.576 11   0.576 -0.887 0.519
## 3 3 1.24   -0.0957     0.319    -0.300 11   0.770 -0.798 0.606

> LSM <- LSmeans(mm, effect='A')

## List of 2
## $ new.fact.lev:List of 1
## ..$ A: chr [1:3] "1" "2" "3"
## $ grid.data    :'data.frame': 3 obs. of  1 variable:
## ..$ A: chr [1:3] "1" "2" "3"

> LSM

##      A      x estimate std.error statistic df p.value      lwr      upr
## 1 1 1.24    0.2895     0.318     0.910 11   0.382 -0.411 0.990
## 2 2 1.24   -0.1841     0.319    -0.576 11   0.576 -0.887 0.519
## 3 3 1.24   -0.0957     0.319    -0.300 11   0.770 -0.798 0.606
```

1.6 Special details

1.6.1 Summary

More details about how the matrix was constructed is provided by the `summary()` method; for example:

```
> summary(LSM)

##      A      x estimate std.error statistic df p.value      lwr      upr
## 1 1 1.24    0.2895     0.318     0.910 11   0.382 -0.411 0.990
## 2 2 1.24   -0.1841     0.319    -0.576 11   0.576 -0.887 0.519
## 3 3 1.24   -0.0957     0.319    -0.300 11   0.770 -0.798 0.606
```

1.6.2 Specification of effect

The **effect** argument can be specified in two ways:

```

> ## 1) a vector of characters and
> LE_matrix(mm, effect= c("A", "C"))

## List of 2
## $ new.fact.lev:List of 2
## ..$ A: chr [1:3] "1" "2" "3"
## ..$ C: chr [1:2] "1" "2"
## $ grid.data : 'data.frame': 6 obs. of 2 variables:
## ..$ A: chr [1:6] "1" "2" "3" "1" ...
## ..$ C: chr [1:6] "1" "1" "1" "2" ...
##      (Intercept) A2 A3      B2      B3 C2      x
## [1,]           1  0  0 0.3333 0.3333  0 1.242
## [2,]           1  1  0 0.3333 0.3333  0 1.242
## [3,]           1  0  1 0.3333 0.3333  0 1.242
## [4,]           1  0  0 0.3333 0.3333  1 1.242
## [5,]           1  1  0 0.3333 0.3333  1 1.242
## [6,]           1  0  1 0.3333 0.3333  1 1.242

> ## 2) a right hand sided formula:
> ## LE_matrix(mm, effect= ~ A + C)

```

1.6.3 Fixing values in the at list

We can fix values of the explanatory variables using the `at=` argument: For example:

```

> LE_matrix(mm, at=list(A=1, x=2))

## List of 2
## $ new.fact.lev:List of 1
## ..$ A: num 1
## $ grid.data : 'data.frame': 1 obs. of 1 variable:
## ..$ A: chr "1"
##      (Intercept) A2 A3      B2      B3 C2 x
## [1,]           1  0  0 0.3333 0.3333 0.5 2

```

1.6.4 Combining effect and at

```

> LE_matrix(mm, effect= ~ A + C, at=list(A=1))

## List of 2
## $ new.fact.lev:List of 2
## ..$ C: chr [1:2] "1" "2"

```

```
## ..$ A: num 1
## $ grid.data :'data.frame': 2 obs. of 2 variables:
## ..$ C: chr [1:2] "1" "2"
## ..$ A: chr [1:2] "1" "1"
##      (Intercept) A2 A3      B2      B3 C2      x
## [1,]           1  0  0 0.3333 0.3333  0 1.242
## [2,]           1  0  0 0.3333 0.3333  1 1.242
```

1.6.5 Total average

Omitting effect as in

```
> LE_matrix(mm)

##      (Intercept)      A2      A3      B2      B3 C2      x
## [1,]           1 0.3333 0.3333 0.3333 0.3333 0.5 1.242

> LSmeans(mm)

## estimate std.error statistic df p.value    lwr    upr
## 1  0.00321    0.184     0.0175 11  0.986 -0.401 0.407
```

gives the “total average”.

1.6.6 Log transformed covariates

Unless otherwise specified, numerical covariates are fixed at their average value, for example for x below:

If we use $\log(x)$ instead we will get an error when calculating LS-means. Instead one must modify data:

```
> mm2 <- lm(y ~ A + B + C + log(x), data=dd)
> ## LSmeans(mm2, effect=~A) ## Will fail
> mm2 <- lm(y ~ A + B + C + log.x,
            data=transform(dd, log.x=log(x)))
> LSmeans(mm2, effect=~A)

## List of 2
## $ new.fact.lev:List of 1
## ..$ A: chr [1:3] "1" "2" "3"
## $ grid.data :'data.frame': 3 obs. of 1 variable:
## ..$ A: chr [1:3] "1" "2" "3"
## A log.x estimate std.error statistic df p.value    lwr    upr
```

```
## 1 1 -0.993    0.315    0.373    0.846 11    0.416 -0.505 1.135
## 2 2 -0.993   -0.254    0.373   -0.682 11    0.509 -1.075 0.566
## 3 3 -0.993   -0.051    0.371   -0.138 11    0.893 -0.868 0.765

> LSmeans(mm2, effect=~A)$L

## List of 2
## $ new.fact.lev:List of 1
## ..$ A: chr [1:3] "1" "2" "3"
## $ grid.data : 'data.frame': 3 obs. of 1 variable:
## ..$ A: chr [1:3] "1" "2" "3"
## NULL
```

This also highlights what is computed: The average of the log of x ; not the log of the average of x .

1.6.7 Powers of covariates

Similar phenomenon for powers of covariates:

```
> mm2 <- lm(y ~ A + B + C + x + I(x^2), data=dd)
> LSmeans(mm2, effect=~A)

## List of 2
## $ new.fact.lev:List of 1
## ..$ A: chr [1:3] "1" "2" "3"
## $ grid.data : 'data.frame': 3 obs. of 1 variable:
## ..$ A: chr [1:3] "1" "2" "3"
## A x I(x^2) estimate std.error statistic df p.value lwr upr
## 1 1 1.24 3.6 -0.331 0.430 -0.770 10 0.459 -1.288 0.626
## 2 2 1.24 3.6 -0.396 0.306 -1.292 10 0.225 -1.079 0.287
## 3 3 1.24 3.6 -0.137 0.286 -0.477 10 0.644 -0.774 0.501

> LSmeans(mm2, effect=~A)$L

## List of 2
## $ new.fact.lev:List of 1
## ..$ A: chr [1:3] "1" "2" "3"
## $ grid.data : 'data.frame': 3 obs. of 1 variable:
## ..$ A: chr [1:3] "1" "2" "3"
## NULL
```

Above $I(x^2)$ is the average of the squared values of x ; not the square of the average of x , cfr. the following.


```

> mm2 <- lm(y ~ A + B + C + x + x2,
            data=transform(dd, x2=x^2))
> LSmeans(mm2, effect=~A)

## List of 2
## $ new.fact.lev:List of 1
## ..$ A: chr [1:3] "1" "2" "3"
## $ grid.data : 'data.frame': 3 obs. of 1 variable:
## ..$ A: chr [1:3] "1" "2" "3"
## A x x2 estimate std.error statistic df p.value lwr upr
## 1 1 1.24 3.6 -0.0398 0.332 -0.120 10 0.907 -0.780 0.700
## 2 2 1.24 3.6 -0.1049 0.289 -0.363 10 0.724 -0.749 0.539
## 3 3 1.24 3.6 0.1544 0.314 0.492 10 0.633 -0.544 0.853

> LSmeans(mm2, effect=~A)$L

## List of 2
## $ new.fact.lev:List of 1
## ..$ A: chr [1:3] "1" "2" "3"
## $ grid.data : 'data.frame': 3 obs. of 1 variable:
## ..$ A: chr [1:3] "1" "2" "3"
## NULL

```

2 Miscellaneous

2.1 Excerpt of the CO2 data

For illustration we use subsets of the CO2 data:

```

> data(CO2)
> C02 <- subset(CO2, conc < 300) ## OK
> C02.bal <- C02
> rownames(C02.bal) <- NULL

```

Data is perfectly balanced. We also use an unbalanced subset of the data

```

> C02.ubal <- C02.bal[-c(1, 5, 12, 17, 18, 19, 20, 28),]
> C02.ubal |> head()

## Plant Type Treatment conc uptake
## 2 Qn1 Quebec nonchilled 175 30.4
## 3 Qn1 Quebec nonchilled 250 34.8
## 4 Qn2 Quebec nonchilled 95 13.6

```

```
## 6   Qn2 Quebec nonchilled 250 37.1
## 7   Qn3 Quebec nonchilled 95 16.2
## 8   Qn3 Quebec nonchilled 175 32.4

> xtabs(~Type + Treatment + conc, data=CO2.ubal) |>
  ftable(row.vars = "Type")

##           Treatment nonchilled          chilled
##           conc           95 175 250           95 175 250
## Type
## Quebec                2   2   3             3   2   1
## Mississippi           2   2   3             2   3   3
```

2.2 Two linear models for CO2 data

```
> library(broom)
> form.add <- uptake ~ conc + Treatment + Type
> form.int <- uptake ~ conc * Treatment + Type
> fm1.bal <- lm(form.add, data=CO2.bal)
> fm1.ubal <- lm(form.add, data=CO2.ubal)
```

2.3 Linear estimates and LSmeans

Common task 1: Consider this task: Estimate the value of the response uptake for each combination of Type and Treatment . This can be obtained, for example, as follows:

```
> LSmeans(fm1.bal, effect=~Type+Treatment, at=list(conc=100))

## List of 2
## $ new.fact.lev:List of 2
## ..$ Type      : chr [1:2] "Quebec" "Mississippi"
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## $ grid.data    : 'data.frame': 4 obs. of 2 variables:
## ..$ Type      : chr [1:4] "Quebec" "Mississippi" "Quebec" "Mississippi"
## ..$ Treatment: chr [1:4] "nonchilled" "nonchilled" "chilled" "chilled"
##           Type Treatment conc estimate std.error statistic df p.value   lwr   upr
## 1   Quebec nonchilled 100    20.30      1.3      15.56 32 1.76e-16 17.64 22.96
## 2 Mississippi nonchilled 100    11.19      1.3       8.58 32 8.40e-10  8.53 13.85
## 3   Quebec   chilled 100    15.33      1.3      11.75 32 3.77e-13 12.68 17.99
## 4 Mississippi   chilled 100     6.22      1.3       4.77 32 3.87e-05  3.57  8.88

> LSmeans(fm1.ubal, effect=~Type+Treatment, at=list(conc=100))
```

```
## List of 2
## $ new.fact.lev:List of 2
## ..$ Type      : chr [1:2] "Quebec" "Mississippi"
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## $ grid.data    : 'data.frame': 4 obs. of  2 variables:
## ..$ Type      : chr [1:4] "Quebec" "Mississippi" "Quebec" "Mississippi"
## ..$ Treatment: chr [1:4] "nonchilled" "nonchilled" "chilled" "chilled"
##      Type Treatment conc estimate std.error statistic df p.value lwr upr
## 1      Quebec nonchilled 100      21.11      1.53      13.79 24 6.72e-13 17.95 24.27
## 2 Mississippi nonchilled 100      11.68      1.66       7.05 24 2.72e-07  8.26 15.10
## 3      Quebec   chilled 100      15.19      1.48      10.26 24 2.96e-10 12.14 18.25
## 4 Mississippi   chilled 100       5.76      1.48       3.88 24 7.11e-04  2.70  8.82
```

Common task 2: Estimate the value of the response uptake when for each level of Treatment. Formally this question can not be answered under the model because the model gives the conditional mean of uptake given conc, Type and Treatment. There is, so to speak, no way of getting rid of Type. There are different workarounds for this situation: We can average over Type and fix conc at its average. Alternatively, we can fit model without Type effect and repeat the steps above.

```
> LSmeans(fm1.bal, effect=~Treatment)
```

```
## List of 2
## $ new.fact.lev:List of 1
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## $ grid.data    : 'data.frame': 2 obs. of  1 variable:
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
##      Treatment conc estimate std.error statistic df p.value lwr upr
## 1 nonchilled 173      23.6      0.885      26.7 32 1.96e-23 21.8 25.4
## 2   chilled 173      18.7      0.885      21.1 32 2.51e-20 16.9 20.5
```

```
> LSmeans(fm1.ubal, effect=~Treatment)
```

```
## List of 2
## $ new.fact.lev:List of 1
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## $ grid.data    : 'data.frame': 2 obs. of  1 variable:
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
##      Treatment conc estimate std.error statistic df p.value lwr upr
## 1 nonchilled 176      24.3      1.03      23.6 24 4.06e-18 22.2 26.4
## 2   chilled 176      18.4      1.04      17.7 24 2.69e-15 16.2 20.5
```

```

> ## Fit model without Type
> fm2.bal <- update(fm1.bal, .~. - Type)
> fm2.ubal <- update(fm1.ubal, .~. - Type)
>
> LSmeans(fm2.bal, effect=~Treatment)

## List of 2
## $ new.fact.lev:List of 1
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## $ grid.data : 'data.frame': 2 obs. of 1 variable:
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## Treatment conc estimate std.error statistic df p.value lwr upr
## 1 nonchilled 173 23.6 1.42 16.6 33 1.32e-17 20.7 26.5
## 2 chilled 173 18.7 1.42 13.1 33 1.17e-14 15.8 21.5

> LSmeans(fm2.ubal, effect=~Treatment)

## List of 2
## $ new.fact.lev:List of 1
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## $ grid.data : 'data.frame': 2 obs. of 1 variable:
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## Treatment conc estimate std.error statistic df p.value lwr upr
## 1 nonchilled 176 24.4 1.66 14.7 25 8.53e-14 21.0 27.8
## 2 chilled 176 17.6 1.66 10.6 25 9.61e-11 14.2 21.0

```

Notice that the predictions (estimates) are identical for the balanced dataset but not for the unbalanced dataset. Notice also that the standard errors of the predictions are somewhat larger when fitting a model without Type effect. This is to be expected as the variation due to Type goes into the residual.

2.4 Generalized linear models

We can calculate LS-means for e.g. a Poisson or a gamma model. Default is that the calculation is calculated on the scale of the linear predictor.

```

> fm.glm <- glm(form.add, family=Gamma, data=CO2.ubal)
> LSmeans(fm.glm, effect=~Treatment, type="link")

## List of 2
## $ new.fact.lev:List of 1
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## $ grid.data : 'data.frame': 2 obs. of 1 variable:
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## Treatment conc estimate std.error statistic p.value lwr upr

```

```
## 1 nonchilled 176 0.0488 0.00287 17.0 6.29e-65 0.0432 0.0544
## 2 chilled 176 0.0599 0.00343 17.5 2.89e-68 0.0532 0.0667
```

2.5 Generalized estimating equations

For gee-type models we get

```
> library(geepack)
> fm.gee <- geeglm(uptake ~ conc + Treatment + Type,
                   id=Plant, family=Gamma, data=CO2.ubal)
> LSmeans(fm.gee, effect=~Treatment)

## List of 2
## $ new.fact.lev:List of 1
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## $ grid.data : 'data.frame': 2 obs. of 1 variable:
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## Treatment conc estimate std.error statistic p.value lwr upr
## 1 nonchilled 176 0.0488 0.00228 21.4 1.50e-101 0.0443 0.0533
## 2 chilled 176 0.0599 0.00382 15.7 1.55e-55 0.0525 0.0674
```

2.6 Linear mixed effects model

For illustration we treat Plant as a random effect:

```
> library(lme4)

## Loading required package: Matrix

> fm.mix <- lmer(uptake ~ conc + Treatment + Type + (1|Plant), data=CO2.ubal)

## boundary (singular) fit: see help('isSingular')

> LSmeans(fm.mix, effect=~Treatment)

## List of 2
## $ new.fact.lev:List of 1
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## $ grid.data : 'data.frame': 2 obs. of 1 variable:
## ..$ Treatment: chr [1:2] "nonchilled" "chilled"
## Treatment conc estimate std.error statistic df p.value lwr upr
## 1 nonchilled 176 24.3 1.05 23.0 7.37 3.89e-08 21.8 26.8
## 2 chilled 176 18.4 1.06 17.3 7.62 2.16e-07 15.9 20.8
```

Notice that the degrees of freedom by default are adjusted using a Kenward–Roger approximation. Unadjusted degrees of freedom are obtained by setting `adjust.df=FALSE`.

2.7 Pairwise comparisons

We will just mention that for certain other linear estimates, the matrix L can be generated automatically using `glht()` from the **multcomp** package. For example, pairwise comparisons of all levels of `dose` can be obtained with

```
> library("multcomp")

## Loading required package: mvtnorm
## Loading required package: survival
##
## Attaching package: 'survival'
## The following object is masked from 'package:boot':
##
##      aml
## Loading required package: TH.data
## Loading required package: MASS
##
## Attaching package: 'MASS'
## The following object is masked from 'package:doBy':
##
##      shoes
##
## Attaching package: 'TH.data'
## The following object is masked from 'package:MASS':
##
##      geyser

> g1 <- glht(fm2.ubal, mcp(Treatment="Tukey"))
> tidy(g1)

## # A tibble: 1 x 7
##   term      contrast      null.value estimate std.error statistic adj.p.value
##   <chr>    <chr>          <dbl>      <dbl>    <dbl>    <dbl>    <dbl>
## 1 Treatment chilled - nonchilled      0      -6.76     2.36     -2.87     0.00829
```

The L matrix is

```

> L <- g1$linfct
> L

##                (Intercept) conc Treatmentchilled
## chilled - nonchilled          0    0                1
## attr(,"type")
## [1] "Tukey"

> lineest(fm2.ubal, L=L)

##                estimate std.error statistic df p.value   lwr   upr
## chilled - nonchilled   -6.76      2.36    -2.87 25 0.00829 -11.6 -1.91

```